

# Lecture 6, Jan 26, 2024

## Estimators for Gaussian RVs

- Consider  $n$  IID measurements  $X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2) = \mathcal{N}(\mu, v)$
- $f_{\mathbf{X}}(\mathbf{x}; \mu, v) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi v}} e^{-(X_i - \mu)^2 / 2v} = \frac{1}{(2\pi v)^{n/2}} e^{-\sum_{i=1}^n \frac{(X_i - \mu)^2}{2v}}$
- Consider the exponent:
 
$$\begin{aligned} \sum_{i=1}^n (X_i - \mu)^2 &= \sum_i (X_i - M_n + M_n - \mu)^2 \\ &= \sum_i (X_i - M_n)^2 + \sum (M_n - \mu)^2 + 2 \sum_i (X_i - M_n)(M_n - \mu) \\ &= nS_n^2 + n(M_n - \mu)^2 \end{aligned}$$
  - $\hat{S}_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - M_n)^2$  is an estimator for the sample variance
- $f_{\mathbf{X}}(\mathbf{x}; \mu, v) = \frac{1}{(2\pi v)^{n/2}} e^{-\frac{nS_n^2}{2v}} e^{-n(M_n - \mu)^2 / 2v}$ 
  - $\log f_{\mathbf{X}}(\mathbf{x}; \mu, v) = -\frac{n}{2} \log(2\pi) - \frac{n}{2} \log v - \frac{nS_n^2}{2v} - \frac{n(M_n - \mu)^2}{2v}$
  - Differentiate wrt  $\mu$ :  $\frac{n}{v}(M_n - \mu) = 0 \implies \hat{\mu}_{\text{ML}} = M_n$
  - Differentiate wrt  $v$ :  $\frac{n}{2v} + \frac{nS_n^2}{2v^2} + \frac{n(M_n - \mu)^2}{2v^2} = 0 \implies \hat{v}_{\text{ML}} = S_n^2$
- Note:  $E[S_n^2] = \frac{1}{n} E \left[ \sum X_i^2 - 2M_n \sum_i X_i + nM_n^2 \right] = E \left[ \frac{1}{n} \sum_i X_i^2 - M_n^2 \right]$ 

$$= (v + \mu^2) - \left( \frac{v}{n} + \mu^2 \right)$$

$$= \frac{n-1}{n} v$$
  - $E[M_n^2] = \text{Var}[M_n] + E[M_n]^2 = \frac{1}{n} v + \mu^2$
  - This is a biased estimator for the variance!
  - For any finite value of  $n$  instead we use  $S_n'^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - M_n)^2$  which is unbiased
  - This applies not just to Gaussians
- Assume the variance is known and the mean has a Gaussian prior; we want to find the MAP estimate
  - Let  $X_i = \Theta + W_i$  where  $W_i$  is IID noise
  - Assume  $E[W_i] = E[W_i | \Theta] = 0$  and  $\text{Var}[W_i] = \text{Var}[X_i | \Theta = \theta] = \sigma_w^2$ , i.e. noise is independent of  $\theta$  and zero-mean, known and fixed variance
  - The prior is  $f_{\Theta}(\theta) = c_1 e^{-\frac{(\theta - \mu)^2}{2\sigma^2}}$
  - The likelihood is  $f_{\mathbf{X}|\Theta}(\mathbf{x}|\theta) = c_2 \prod_{i=1}^n e^{-\frac{(x_i - \theta)^2}{2\sigma_w^2}}$ 
    - \* Knowing  $\theta$  just gives us the mean of the distribution
    - \* Note the variance that appears here is different than in the prior!
  - The posterior distribution:  $\propto f_{\mathbf{X}|\Theta}(\mathbf{x}|\theta) f_{\Theta}(\theta) = c_3 \exp \left( -\frac{1}{2\sigma_w^2} \sum_{i=1}^n (x_i - \theta)^2 - \frac{1}{2\sigma^2} (\theta - \mu)^2 \right)$
  - The exponent becomes  $\theta^2 \left( \sum_{i=1}^n \frac{1}{2\sigma_w^2} + \frac{1}{2\sigma^2} \right) - 2\theta \left( \sum_{i=1}^n \frac{x_i}{2\sigma_w^2} + \frac{\mu}{2\sigma^2} \right) + c_4$
  - Completing the square:  $\frac{n\sigma^2 + \sigma_w^2}{2\sigma^2 \sigma_w^2} \left( \theta - \frac{\sigma^2 \sigma_w^2}{n\sigma^2 + \sigma_w^2} \left( \frac{nM_n}{\sigma_w^2} + \frac{\mu}{\sigma^2} \right) \right)^2$

- \* This shows that  $\theta$  is also a Gaussian with mean  $\frac{\sigma^2\sigma_w^2}{n\sigma^2 + \sigma_w^2} \left( \frac{nM_n}{\sigma_w^2} + \frac{\mu}{\sigma^2} \right)$  and variance  $\frac{\sigma^2\sigma_w^2}{n\sigma^2 + \sigma_w^2}$
- Since this is a Gaussian we know the maximum occurs at the expectation value
- The MAP estimate is then  $E[\Theta|\mathbf{X}] = \frac{n\sigma^2}{n\sigma^2 + \sigma_w^2} M_n + \frac{\sigma_w^2}{n\sigma^2 + \sigma_w^2} \mu$ 
  - \* As  $n \rightarrow \infty$  the first weight approaches 1, the second approaches zero
  - \* This means as we take more samples, the MAP estimate approaches the ML estimate, as the information from the measurements becomes more important than the prior
- $\text{Var}[\Theta|\mathbf{X}] = \frac{\sigma^2\sigma_w^2}{n\sigma^2 + \sigma_w^2}$ 
  - \* Notice that this goes to zero as  $n \rightarrow \infty$
- In this case,  $\hat{\Theta}_{\text{LMS}} = \hat{\Theta}_{\text{MAP}} = E[\Theta|\mathbf{X}]$