

Lecture 14, Feb 10, 2023

The Normal (Gaussian) Distribution

Definition

Given a mean μ and variance σ^2 , the *normal distribution* is given by

$$n(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, x \in \mathbb{R}$$

- This gives a symmetric bell centered around the mean μ , with width proportional to σ
 - μ is a translation
 - σ gets larger as the curve gets wider and flatter
 - $\lim_{\sigma \rightarrow 0} n(x; \mu, \sigma) = \delta(x - \mu)$
- The Gaussian is important due to the central limit theorem: taking a large number of random variables and taking their average, it will give the normal distribution regardless of the distribution of the individual random variables

$$\begin{aligned} \left(\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} dx \right)^2 &= \frac{1}{2\pi\sigma^2} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2\sigma^2}} dx \int_{-\infty}^{\infty} e^{-\frac{y^2}{2\sigma^2}} dy \\ &= \frac{1}{2\pi\sigma^2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-\frac{x^2+y^2}{2\sigma^2}} dx dy \\ &= \frac{1}{2\pi\sigma^2} \int_0^{2\pi} \int_0^{\infty} e^{-\frac{r^2}{2\sigma^2}} r dr d\theta \\ &= \frac{1}{4\pi\sigma^2} \int_0^{2\pi} \int_0^{\infty} e^{-\frac{s}{2\sigma^2}} ds d\theta \\ &= \frac{1}{2\sigma^2} \int_0^{\infty} e^{-\frac{s}{2\sigma^2}} ds \\ &= \frac{1}{2\sigma^2} \left[-2\sigma^2 e^{-\frac{s}{2\sigma^2}} \right]_0^{\infty} \\ &= 1 \end{aligned}$$

- $E[X] = \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$
 - $= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} (\sigma z + \mu) e^{-\frac{z^2}{2}} \sigma dz$
 - $= \frac{\sigma}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z e^{-\frac{z^2}{2}} dz + \frac{\mu}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} dz$
 - $= 0 + \mu \int_{-\infty}^{\infty} n(z; 0, 1) dz$
 - $= \mu$
 - Substitute $z = \frac{x - \mu}{\sigma}$
 - $\int_{-\infty}^{\infty} z e^{-\frac{z^2}{2}} dz = 0$ because the integrand is an odd function
- Using a similar argument we may show that the variance is σ^2

Definition

$n(x; 0, 1)$ is referred to as the *standard normal distribution*

$$\Phi(x) = \int_{-\infty}^x n(y; 0, 1) dy$$

is the cumulative distribution function of the standard normal, so

$$P(A \leq X \leq B) = \Phi(B) - \Phi(A)$$

- Note Φ is not analytically evaluable, so there are usually tables of values for it
- Suppose X has PDF $n(x; \mu, \sigma)$; let $Z = \frac{X - \mu}{\sigma}$, then Z has PDF $n(x; 0, 1)$, which is the standard normal

$$\begin{aligned} - P(X \leq x) &= \int_{-\infty}^x n(x; \mu, \sigma) dx \\ &= \int_{-\infty}^x \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(t-\mu)^2}{\sigma^2}} dt \\ &= \int_{-\infty}^{\frac{x-\mu}{\sigma}} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{s^2}{2}} \sigma ds \\ &= \int_{-\infty}^{\frac{x-\mu}{\sigma}} n(s; 0, 1) ds \\ &= P\left(Z \leq \frac{x-\mu}{\sigma}\right) \\ &= \Phi\left(\frac{x-\mu}{\sigma}\right) \end{aligned}$$

- Therefore $P(A \leq X \leq B) = \Phi\left(\frac{A-\mu}{\sigma}\right) - \Phi\left(\frac{B-\mu}{\sigma}\right)$
- Example: Suppose X is a random variable with distribution $n(x; 5, 2)$; find $P(-1 \leq X \leq 4)$
 - Need to transform this into the standard normal
 - Let $Z = \frac{X-5}{2}$ then Z has the standard normal distribution and CDF Φ
 - $P(-1 \leq X \leq 4) = P\left(\frac{-1-5}{2} \leq \frac{X-5}{2} \leq \frac{4-5}{2}\right) = P\left(-3 \leq Z \leq -\frac{1}{2}\right) = \Phi\left(-\frac{1}{2}\right) - \Phi(-3) = 0.3072$
 - $\Phi(x)$ is `normcdf` in MATLAB